Probability thresholds estimates for coloring properties of random hypergraphs

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joint work with D. Shabanov

Probabilistic Combinatorics Online

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Independent set

This is a 4-uniform hypergraph H on 8 vertices with 6 edges (surfaces of the cube):

What is the maximum number of vertices we can take such that there's no edge from which we took all 4 of them?



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The answer is 6, so is the independence number of this hypergraph.

j-independent set

For the same hypergraph what is the maximum number of vertices we can take such that there's no edge from which we took more than 2 vertices? More than 1?

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j-independent set

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2-independence number of this hypergraph is 4.

1-independence number of this hypergraph is 2.

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j-chromatic numbers

How many colors are needed to color vertices of hypergraph such that every color is j-independent set?



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Definitions

- For an integer j, a j-independent set in a hypergraph H(V, E) is a subset A ⊂ V such that for every edge e ∈ E : |e ∩ A| ≤ j;
- A coloring of hypergraph H = (V, E) is called *j*-proper if every color class is *j*-independent;
- The *j*-chromatic number \(\chi_j(H)\) of hypergraph \(H = (V, E)\) is the minimal number of colors in a *j*-proper coloring of \(H;\)
- ▶ In the case of *k*-uniform hypergraph $\chi_{k-1}(H) = \chi(H)$;
- All chromatic numbers $\chi_j(H)$ with $j \ge k/2$ are called **weak**.

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Let's consider binomial model H(n, k, p) of a random hypergraph.

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Let's consider binomial model H(n, k, p) of a random hypergraph.



Constant probability

Let's consider j and k are fixed while n tends to infinity. Also let's assume j < k - 1 and $p \in (0, 1)$ is some constant.

- Let X be the number of subsets of j + 1 vertices that are not contained in any edge of hypergraph H(n, k, p).
- Probability that a set of j + 1 vertices is not contained in any edge of hypergraph H(n, k, p) equals

$$(1-p)^{\binom{n-j-1}{k-j-1}}$$

Therefore,

$$\mathsf{E} X = \binom{n}{j+1} (1-\rho)^{\binom{n-j-1}{k-j-1}} \to 0.$$

• Then $\chi_j(H(n, k, p)) = \lceil \frac{n}{j} \rceil$ asymptotically almost surely (a.a.s.).

Dense case

Let's take $1 \leq j \leq k - 1$. Denote

$$d^{(j)} = d^{(j)}(n,p) = j \binom{k-1}{j} \binom{n-1}{k-1} p.$$

- Shamir, 1989: $d^{(j)} > n^{j-\varepsilon}, \varepsilon > 0$ and $d^{(j)} = o(n^j \ln n);$
- **•** Krivelevich, Sudakov, 1998: $d^{(j)} = o(n^j)$ and $d^{(j)} \to \infty$;

$$\chi_j(H(n,k,p)) \underset{\mathsf{P}}{\sim} \left(\frac{d^{(j)}}{(j+1) \ln d^{(j)}} \right)^{1/j}, \ n \to \infty$$

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Sparse case, arbitrary r, j = k - 1

Let's now consider the case $p = cn/\binom{n}{k}$.

- ► Theorem (Dyer, Frieze, Greenhill, 2015): If $c > r^{k-1} \ln r - \frac{\ln r}{2}$ then a.a.s. $\chi(H(n, k, p)) > r$.
- ► Theorem (Dyer, Frieze, Greenhill, 2015): If $c < r^{k-1} \ln r - \frac{r-1}{r} (1 + \ln r) + O(\frac{k^2 \ln r}{r^{k-1}})$ then a.a.s. $\chi(H(n, k, p)) \leq r$.
- ▶ Theorem (Ayre, Coja-Oghlan, Greenhill, 2015): For $r > r_0(k)$, if $c < r^{k-1} \ln r - \frac{\ln r}{2} - \ln 2 + o_r(1)$ then a.a.s. $\chi(H(n, k, p)) \leq r$.
- Theorem (Shabanov, 2017): For $k \ge 4$, if $c < r^{k-1} \ln r - \frac{\ln r}{2} - \frac{r-1}{r} + o_{r,k}(1)$ then a.a.s. $\chi(H(n,k,p)) \le r.$

Sparse case, r = 2, j = k - 1

• Theorem (Coja-Oghlan, Panagiotou, 2012): There exists $\varepsilon_k = 2^{-k(1+o_k(1))}$, such that if

$$c < 2^{k-1}\ln 2 - \frac{\ln 2}{2} - \frac{1}{4} - \varepsilon_k$$

then a.a.s. $\chi(H(n,k,p)) \leq 2$. On the other hand, if

$$c > 2^{k-1} \ln 2 - \frac{\ln 2}{2} - \frac{1}{4} + \varepsilon_k$$

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then a.a.s. $\chi(H(n, k, p)) > 2$.

Sparse case,
$$r = 2$$
, $k - j = o(\sqrt{k})$

• Theorem (Semenov, 2017): There exist C_l , C_u and k_0 , such that if $k > k_0$, $2 \le k - j = o(\sqrt{k})$ and

$$c > rac{2^{k-1} \ln 2}{\sum_{s=0}^{k-j-1} {k \choose s}} - rac{\ln 2}{2} + C_u \cdot k^{k-j-1} \cdot 2^{-k},$$

then a.a.s. $\chi_j(H(n,k,p)) > 2$. On the other hand, if

$$c < rac{2^{k-1} \ln 2}{\sum_{s=0}^{k-j-1} {k \choose s}} - rac{\ln 2}{2} - C_l \cdot (k/8)^{j+1-k},$$

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then a.a.s. $\chi_j(H(n,k,p)) \leq 2$.

Sparse case, arbitrary r, $k - j = o(k^{1/4})$

▶ Theorem (Semenov, Shabanov, 2020): For any r > 2, there exist C_l , C_u and $k_0 = k_0(r)$, such that if $k > k_0$, $2 \le k - j = o(k^{1/4})$ and

$$c > \frac{r^{k-1} \ln r}{\sum_{s=0}^{k-j-1} {\binom{k}{s}} (r-1)^s} - \frac{\ln r}{2} + C_u \cdot {\binom{k}{j+1}} \cdot r^{-j},$$

then a.a.s. $\chi_j(H(n,k,p)) > r$. On the other hand, if

$$c < rac{r^{k-1} \ln r}{\sum_{s=0}^{k-j-1} {k \choose s} (r-1)^s} - rac{\ln r}{2} - C_l \cdot k^{j-k+1},$$

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then a.a.s. $\chi_j(H(n,k,p)) \leq r$.

Proof (upper bound)

▶ Different model H₁(n, k, m): we randomly choose m = ⌈cn⌉ edges with replacement from the set of all possible edges (some edges may repeat); For c' > c and p' = c'n/ \binom{n}{k}

 $\mathsf{P}(\chi_j(H_1(n,k,\lceil cn\rceil))>r)\leqslant \mathsf{P}(\chi_j(H(n,k,p')>r)+o_n(1).$

- Counting proper colorings: let X_n be a random variable corresponding to the number of *j*-proper colorings of H₁(n, k, [cn]);
- First moment method:

 $\mathsf{P}(X_n > 0) \leqslant \mathsf{E}X_n.$

Proof (lower bound)

Different model H₂(n, k, m): we choose m = ⌈cn⌉ independent edges and in every edge choose k vertices also randomly and independently (some edges may repeat and be non-proper); For c' < c and p' = c'n/ {n \atop k}</p>

 $\mathsf{P}(\chi_j(H_2(n,k,\lceil cn\rceil))\leqslant r)\leqslant \mathsf{P}(\chi_j(H(n,k,p')\leqslant r)+o_n(1).$

Sharp threshold: Based on the result of Hatami and Molloy, there exists a function p̂ = p̂(n) such that for every ε > 0,

$$\mathsf{P}(\chi_j(\mathsf{H}(n,k,p))\leqslant r)
ightarrow egin{cases} 1, & p<(1-arepsilon)\hat{
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Hence, we need to show that $P(\chi_j(H_2(n, k, m) \leq r)$ is bounded from zero.

Second moment method

- We can consider only the case of n divisible by r.
- Counting proper balanced colorings: let X_n be a random variable corresponding to the number of *j*-proper balanced colorings of H₂(n, k, [cn]), i.e. color classes are of size n/r;

$$\mathsf{P}(\chi_j(H_2(n,k,\lceil cn\rceil))\leqslant r) \geqslant \mathsf{P}(X_n>0);$$

Second moment method: The Paley-Zygmund inequality says that

$$\mathsf{P}(X_n > 0) \geqslant \frac{(\mathsf{E}X_n)^2}{\mathsf{E}X_n^2};$$

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So, the final step is to show that $EX_n^2 = O_k((EX_n)^2)$.

Second moment calculation

► To calculate the second moment we use matrices A ∈ A of size r by r with the property:

$$\sum_{i=1}^{r} a_{iu} = n/r, \sum_{u=1}^{r} a_{iu} = n/r.$$

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• Let's also denote $J_r \in A$ as a matrix with entries that are all the same and equal n/r^2 .

Helper functions

Then, we introduce functions:

$$\mathcal{H}(A) = -\sum_{i,u=1}^{r} \frac{a_{iu}}{n} \ln \frac{ra_{iu}}{n};$$
$$\mathcal{E}(A) = \ln\left(1 - 2r^{1-k}\sum_{s=j+1}^{k} \binom{k}{s}(r-1)^{k-s} + \sum_{i,u=1}^{r}\sum_{s=j+1}^{k} \binom{k}{s}\sum_{h=0}^{k-s}\sum_{t=0}^{s-j-1+h} \binom{k-s}{h}\binom{s}{t}.$$
$$\left(\frac{\frac{n}{r} - a_{iu}}{n}\right)^{h+t} \left(\frac{a_{iu}}{n}\right)^{s-t} \left(\frac{\frac{n(r-2)}{r} + a_{iu}}{n}\right)^{k-h-s}\right).$$

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Crucial technical lemma

For c that satisfies conditions of the theorem, the expression $\mathcal{G}_c(A) = \mathcal{H}(A) + c \cdot \mathcal{E}(A)$ takes it's minimum value when $A = J_r$. Which comes from the following

Lemma: There exist b = b(k, r) > 0, such that for every $A = (a_{iu}, i, u = 1, ..., r)$ from A

$$\mathcal{G}_{c}(A) - \mathcal{G}_{c}(J_{r}) \geq b \sum_{i,u=1}^{r} \left(\frac{a_{iu}}{n} - \frac{1}{r^{2}}\right)^{2}$$